Measures of Precision





Overview

- How to quantify uncertainty
- Why variance is important
- Components of variation in distance sampling
- Controlling variance
- Estimating variance
- Analytic
- Bootstrap
- Confidence Intervals





Consider an artificial population

 $D = 500 \text{ per unit}^2 \text{ (no density gradient)}$

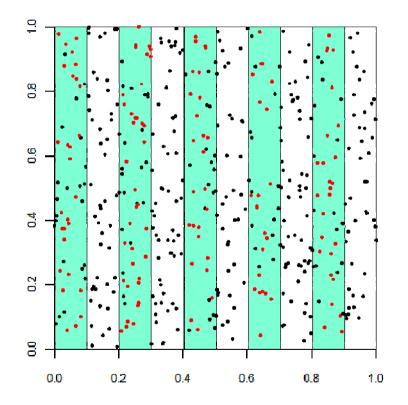
Design: 5 transects equally-spaced (w=0.05)

Results:

$$n = 140$$

$$\hat{f}(0) = 34.6$$

$$\hat{D} = 484.4$$







Consider a duplicate survey

Same population model

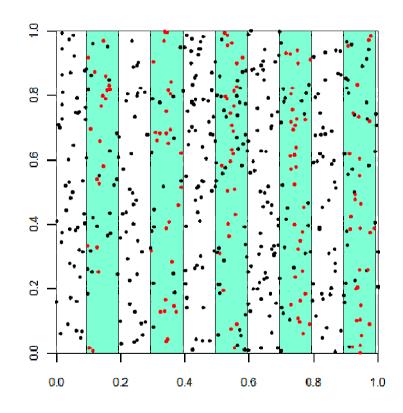
Same survey design (with a new random start point)

Results:

$$n = 139$$

$$\hat{f}(0) = 37.6$$

$$\hat{D} = 522.1$$







Imagine repeating this process over and over, using the same survey design and a population drawn from the same density model

Each survey will yield:

A different value for n

A different value for $\hat{f}(0)$

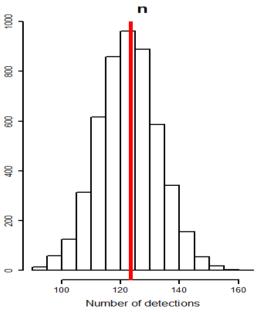
A different value for \hat{D}

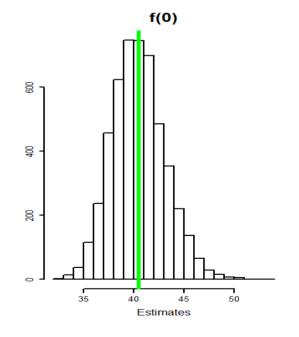


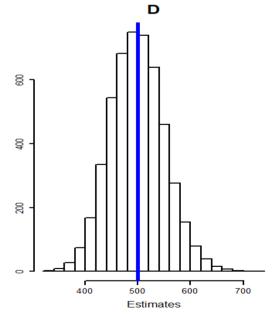


What happens if we repeat this simulated survey 10,000 times?

We end up with **distributions** for n, $\hat{f}(0)$ and \hat{D}











We are interested in the **hypothetical long-run** behaviour of our estimator

$$\hat{D} = \frac{n\hat{f}(0)}{2L}$$

How variable are the estimates?

E.g. what is the variance of the distribution for \hat{D} ?

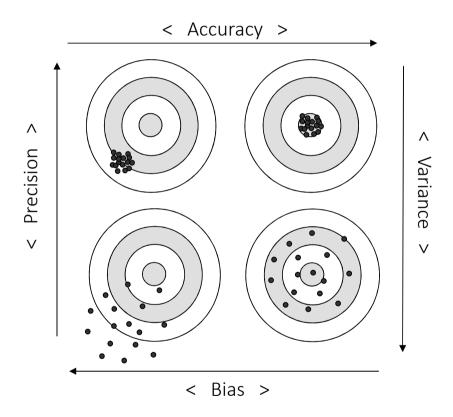
What is the average value of the estimates?

E.g. is the distribution for \hat{D} centred on the truth?





Bias vs. Variance





Low precision = high variance = high uncertainty



Quantifying uncertainty

Different ways of measuring uncertainty:

Variance = the average squared difference from the mean (the inverse of precision)
 If the estimator for D is unbiased, then

$$Var[\hat{D}] = E[(\hat{D} - D)^2]$$

2. Standard error = the standard deviation of an estimator (i.e. the square root of estimator variance)

$$Se[\hat{D}] = \sqrt{Var[\hat{D}]}$$





Quantifying uncertainty

3. Coefficient of Variation (CV) = the standard error dived by the mean (i.e. a standardised version of the standard error) $C = \hat{C}$

$$CV[\hat{D}] = \frac{Se[\hat{D}]}{E[\hat{D}]}$$

Useful for comparing variances when the scale and/or the units of measurement differ

E.g. consider two variables: X has mean = 100 and variance = 400, Y has mean = 1 and variance = 0.04

$$CV[X] = \frac{\sqrt{400}}{100} = \frac{20}{100} = 0.2 = 20\%$$
 $CV[Y] = \frac{\sqrt{0.04}}{1} = \frac{20}{100} = 0.2 = 20\%$





Quantifying uncertainty

4. Confidence Interval (CI) = a range of plausible values for the truth

Calculations are based on variance

Different ways to calculate CIs, depending on the data, e.g.

Normal

Lognormal (available in Distance)

Bootstrap (available in Distance)

More about CIs later...





Why is variance important?

- •In a real survey, we use an estimator and the survey data to produce a single estimate for *D*
- •If the estimator variance is low, then individual estimates are more likely to be close to the truth (assuming low bias)
- •If estimator variance is high, then individual estimates are more likely to be far from the truth
- •For reliable results, we want estimators with LOW variance (and low bias!)





We can break down the familiar distance sampling density estimator (for line transects with no clusters) into three components:

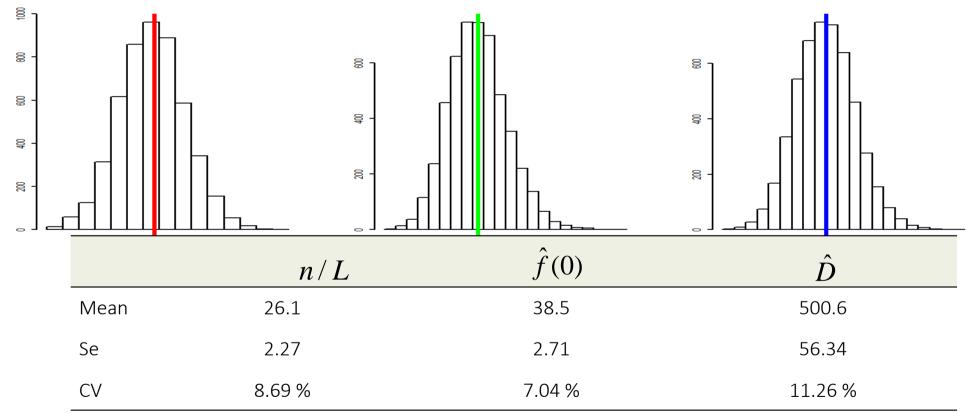
$$\hat{D} = \frac{n\hat{f}(0)}{2L} = \frac{1}{2} \times \frac{n}{L} \times \hat{f}(0)$$
Constant (no variance)

Encounter rate

University of St Andrews



We can calculate variance measures separately for each component







- ullet The variance of \hat{D} is affected by the variance of its components
- •If the variance of n is high, then the variance of n/L will be high and the variance of \hat{D} will be high
- •Similarly, if the variance of $\hat{f}(0)$ is high then the variance of \hat{D} will be high
- •So for reliable estimates, we want Var[n/L] and $Var[\hat{f}(0)]$ to be low





Distance provides several variance measures for each component

Parameter	Point Estimate	Standard Error	Percent Coef. of Variation	95% Pe Confidenc	rcent e Interval
f(0)	1.5726	0.19139	12.17	1.2304	2.0102
р	0.42391	0.51589E-01	12.17	0.33165	0.54185
ESW	0.63587	0.77383E-01	12.17	0.49747	0.81277
n/L	0.80579E-01	0.25990E-01	32.25	0.40601E-01	0.15992
DS	0.63361E-01	0.21843E-01	34.47	0.31088E-01	0.12914
E(S)	2.0143	0.20292	10.07	1.6433	2.4692
D	0.12763	0.45838E-01	35.92	0.61445E-01	0.26510
N	10849.	3896.4	35.92	5223.0	22534.





Encounter rate variance

The encounter rate = n/L = the number of detections per unit of distance

The variance of n/L is related to the variance of n, and therefore to the variances of counts for individual transects

Therefore, if counts from individual transects are highly variable the variance of n/L will also be high $Var[n] = Var[n_1] + ... + Var[n_k]$ assumes independence





- We can use this knowledge of encounter rate variance to help design good surveys
- •Three main ways we can reduce encounter rate variance:
- Use systematic survey designs
- Run transects parallel to density gradients
- Use designs with several transects

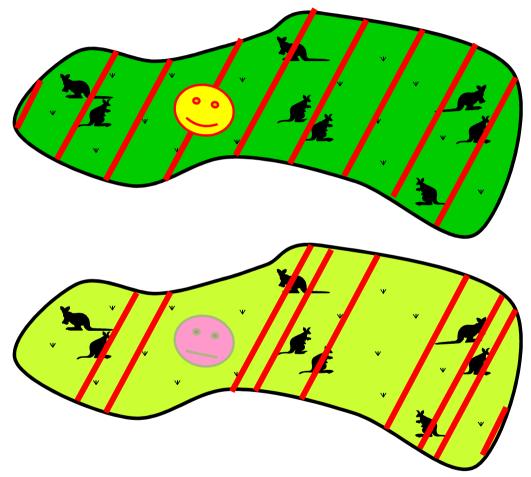




1. Use systematic survey designs

These give lower variance than completely random designs

More likely to give even coverage of the survey region



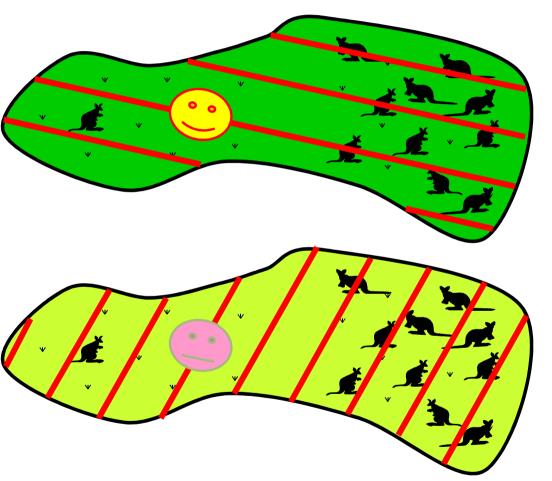




2. Run transects parallel to known density gradients

i.e. perpendicular to gradient contours

Lines are more likely to have similar encounter rates since each line will cover the full range in density



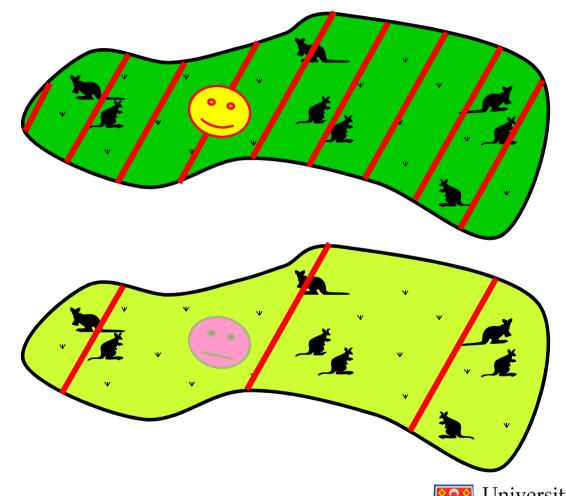




3. Use many lines (or points)

This will give good spatial coverage

Using longer lines will also help to reduce variation between lines (because each transect will have a larger sample size)







Recap

- Density estimates will vary between (hypothetical) duplicate surveys
- Density estimators with low bias and low variance are more reliable, since individual
 estimates are more likely to be close to the truth
- ullet High variance in the encounter rate and the fitted detection function will lead to high variance in \hat{D}
- ullet Good survey design can help to lower the variance of \hat{D} by reducing encounter rate variance





Estimating variance

- •So how do we **measure** the variance?
- •In reality we rarely know the true variance of our estimator, and we can't carry out 1000s of real surveys...
- What we need to do is estimate the variance using the data from a single survey
- •Two methods:
- Analytic
- Bootstrap





We can describe the relationship between the variance of \hat{D} and the variance of its components more formally using a useful approximation known as the <code>Delta method</code>

$$\left\{cv(\hat{D})\right\}^{2} = \left\{cv\left(\frac{n}{L}\right)\right\}^{2} + \left\{cv(\hat{f}(0))\right\}^{2}$$

Rule: when two or more components are multiplied together, squared CVs add





We can check this approximation works using the results of our simulation,

$$\left\{cv(\hat{D})\right\}^2 = 0.1125^2 = 0.01266$$

$$\left\{cv\left(\frac{n}{L}\right)\right\}^2 + \left\{cv(\hat{f}(0))\right\}^2 = 0.0869^2 + 0.0704^2 = 0.01251$$

We can rearrange the squared CV to get an estimate of the variance

$$\operatorname{var}(\hat{D}) \approx \left\{ \hat{D} \times \sqrt{\left\{ cv(\hat{D}) \right\}^2} \right\}^2$$





- ullet To estimate the variance of \hat{D} we need to estimate the squared CVs of the components
- We therefore need estimates for the variances of n/L and $\hat{f}(0)$
- The variance of $\hat{f}(0)$ can be estimated using standard methods (because its parameters are obtained via maximum likelihood)
- But n/L is part of the data, not an estimate, and estimating its variance is more difficult...





- To estimate var(n/L) we need to use data from the individual lines (or points)
- A minimum of 20 replicate lines (or points) is recommended for obtaining a reliable estimate of encounter rate variance
- The (improved) formula used in Distance:

$$\left\{ CV\left(\frac{n}{L}\right) \right\}^{2} = \frac{k}{n^{2}(k-1)} \sum_{i=1}^{k} \ell_{i}^{2} \left(\frac{n_{i}}{\ell_{i}} - \frac{n}{L}\right)^{2}$$

$$k = \text{number} \quad \text{for line i} \quad \text{for line i}$$





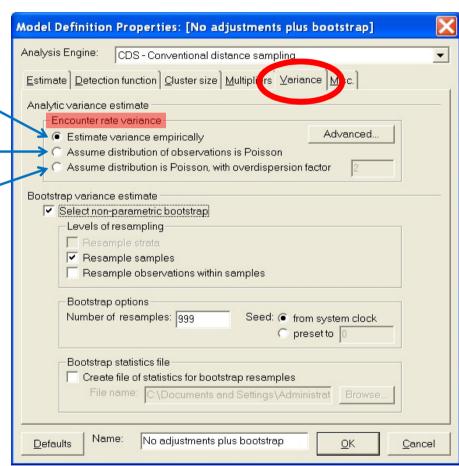
- Note that the formula used in Distance assumes independently arranged lines
- It therefore tends to overestimate encounter rate variance for systematic designs (for which there are no analytic variance estimators)
- Stratification-based methods can help to remove this bias (see later lecture)
- Highlights the difference between the true (and unknown) variance and our ability to estimate it





Three options in Distance:

- 1. Use the formula we have just seen
- 2. Assume var(n)=n (last resort)
- 3. Assume $var(n) = \vartheta n$ (slightly better last resort)







Point Parameter Estimate		Standard Error	Percent Coef. of Variation	95% Percent Confidence Interval		
	ESCIMACE	FILOI	or variation	Confidence		
f(0)	1.5726	0.19139	12.17	1.2304	2.0102	
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			44			

Measurement Units

Density: Numbers/Sq. nautical mi

ESW: nautical miles

Similarly, N and D always have the

Component Percentages of Var(D) same CV

Detection probability : 11.5 Encounter rate : 80.7 CRFFM Cluster size

p and ESW are derived from 1/f(0), so these three share the same CV





To find the **relative contributions** of each component we take the ratio of squared CVs

E.g.
$$100\% \times \frac{\left\{cv(\hat{f}(0))\right\}^{2}}{\left\{cv(\hat{D})\right\}^{2}} = \text{The percentage relative contribution made by f(0)}$$

	Typical values				
Component	Line	Point			
Encounter rate	70-80%	40-50%			
Detection function	<30%	>50%			



Parameter	Point Estimate		Percent Coef. of Variation		
	1.5726			1.2304	2.0102
p	0.42391	0.51589E-01	12.17	0.33165	0.54185
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	Units				
ESW: na	umbers/Sq. nau uutical miles Percentages of		12. 2	17²/35.92² x	100%





The bootstrap method works as follows:

- 1. Generate a new sample by repeatedly sampling from the original sample randomly and with replacement
 - some units from the original sample may appear more than once in the new sample (others might not appear at all)
 - each new sample must be the same size as the original sample
- 2. Calculate a density estimate using the new sample
- 3. Repeat steps 1 & 2 a large number of times (e.g. 999) to obtain multiple density estimates





- Works well if the original sample is large and representative
- The distribution of density estimates approximates the true distribution that we would (theoretically) get from duplicate surveys
- The variance of the bootstrap estimates can be used as an estimate of the true variance
- In Distance we resample the individual transects





- For example, consider a survey with 12 replicate lines
 - Bootstrap sample 1:
 - Transects: 5, 12, 1, 7, 6, 11, 7, 6, 9, 7, 11, 2
 - Density estimate = D_1
 - Bootstrap sample 2:
 - Transects: 3, 4, 9, 1, 12, 7, 8, 11, 1, 3, 2, 12
 - Density estimate = D_2
- Do this B times and use the variance of the B density estimates as an estimate of $var(\hat{D})$

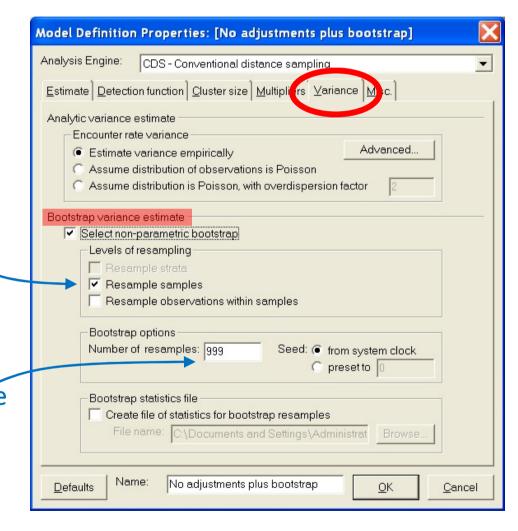




The usual option (samples = transects)

The number of bootstrap samples to use (test on a small number first to ensure all is properly set up)







Confidence Intervals

- Confidence intervals (CIs) give us a range of plausible values for the truth
- Constructed using data from a single sample
- If we were to carry out multiple surveys and construct 95% CIs from each survey, we would expect 95% of those CIs to contain the true value
- To calculate CIs we need to know the **shape** of the distribution of estimates





Confidence Intervals - Analytic

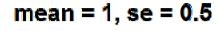
• Two choices:

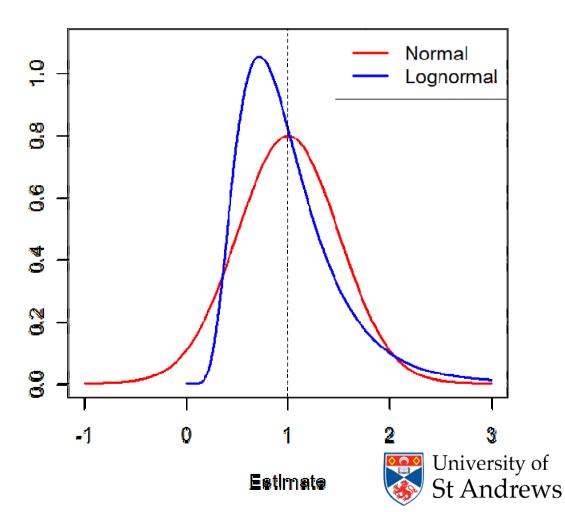
Normal

- symmetrical
- easy to use
- allows negative values

Lognormal

- asymmetric (skewed)
- trickier to use
- typically higher interval limits
- does not allows negative values







Confidence Intervals - Analytic

Distance uses 95% lognormal Cls

Parameter	Point Estimate	Standard Error	Percent Coef. of Variation	95% Pe Confidence	rcent e Interval
f(0)	1.5726	0.19139	12.17	1.2304	2.0102
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$$\left(\frac{\hat{D}}{C}, \hat{D} \times C\right)$$
 $C = \exp\left[1.96\sqrt{\ln\left(1 + \left(cv(\hat{D})\right)^2\right)}\right]$





Confidence Intervals – Bootstrap

We can use the bootstrap estimates to construct CIs for the true density in two ways:

Parametric

Use the lognormal CI method with the bootstrap estimate of variance instead of the analytic estimate

Non-parametric

Place the bootstrap estimates in order of increasing size and use percentiles as the CI limits (e.g. for a 95% CI using 999 bootstrap estimates, take the 25th estimate as the lower limit and the 975th estimate as the upper limit)





Confidence Intervals – Bootstrap

Both options are provided in Distance

Pooled	Estimates:						
		Estimate	%CV	#	df	95% Confider	nce Interval
	DS	0.20906E-01	 38.14	999	19.43	0.96777E-02	0.45162E-01
						0.12166E-01	0.41557E-01
	D	0.39125E-01	42.24	999	23.93	0.16953E-01	0.90294E-01
						0.22020E-01	0.81797E-01
	N	27987.	42.24	999	23.9	12127.	64589.
					(15752.	58511.
						Option 1 (parar	netric) \
Note:	Confidence inte	rval 1 uses 1	bootstr	ap SE	and log-	normal 95% in	ntervals.
	Interval 2 is t	he 2.5% and	97.5%	quant	iles of t	he bootstrap	estimates.

Option 2 (non-parametric)





Other advantages of the bootstrap

Ambivalent model fit

E.g. different detection functions can be fitted to each bootstrap resample if it is difficult to choose between competing models (model uncertainty is therefore incorporated into the bootstrap density estimates)

Checking independence of components

If the bootstrap and analytical CIs are very different, then the 'squared CVs add' rule may have been violated (i.e. due to non-independence of the components of the density estimator – a necessary assumption for this approximation to work well)





Further reading

- Section 3.6 of Buckland et al. (2001) Introduction to Distance Sampling
- Sections 6.3.1.2 (lines) and 6.3.2.2 (points) of Buckland et al. (2015) Distance Sampling: Methods and Applications
- Fewster et al. (2009) Estimating the encounter rate variance in distance sampling. Biometrics 65: 225-236.



